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### Introduction and motivation

Large systems Modelling

#### Updating beliefs

The Bayes linear approach Exchangeable events Making decisions

### Application: corrosion monitoring

Corrosion monitoring Data characteristics Bayes linear variance learning Model diagnostics

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### Conclusions and future work

### Large systems

- Research: galaxy evolution, climate change
- Manufacturing: fouling, corrosion, fatigue
- Environmental: ground, water and airborne monitoring
- Commerce: financial, transactional, software







Large systems

## System characteristics

- High dimensional (> 1000 variables)
- Dependent variables (e.g. in time or space)
- Evolves (e.g. in time)
- Observed with error
- Observing complete system prohibitively costly

Introduction and motivation

Large systems

# Method components

- 1. Specify model
  - Partial belief structure
  - Exchangeability assumptions (if any)
- 2. Simulate to estimate full belief structure
- 3. Adjust expectations given beliefs and observations
  - Incomplete and irregular observations
  - ► Learn about system level and (co-)variance structure

- 4. Simulate adjusted system to forecast
- 5. Make decision
  - Expected loss to optimise decision

Learning about large industrial systems
Introduction and motivation
Modelling

### Typical model specification

- Two spatial dimensions (I, c), one temporal (t)
- Observations in time (t) and one spatial dimension (c) only

- Observations with error (eylct)
- **Global** evolution  $(\epsilon_{\Theta ct})$  with respect to t and c
- ▶ Local evolution in *I* dimension  $(\epsilon_{rlct})$  relative to global

# Typical model form

Observation:	$Y_{ct} = f_l \left( Z_{lct} + \epsilon_{Ylct} \right)$	$\operatorname{Var}(\epsilon_{Y ct}) = \sigma_Y^2$
System:	$Z_{lct} = \mathbf{F} \mathbf{\Theta_{ct}} + r_{lct}$	
Global Effects:	$\boldsymbol{\Theta}_{ct} = \mathbf{G} \boldsymbol{\Theta}_{ct-1} + \boldsymbol{\epsilon}_{\boldsymbol{\Theta}ct}$	$\operatorname{Var}(\epsilon_{\Theta t}) = \Sigma_{\Theta}$
Local Effects:	$r_{lct} = g(r_{lct-1}) + \epsilon_{rlct}$	$\operatorname{Var}(\epsilon_{rlct}) = \sigma_{rl}^2$

- f<sub>l</sub> reduces (or "integrates" over) l
- g describes local evolution
- F and G are regression and system evolution matrices

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Introduction and motivation
Modelling

# Partial to full beliefs

Specify partial beliefs:

- ▶ Specify model form *f*<sub>*l*</sub>, **F**, **G** and *g*
- Specify variance structures  $\sigma_Y^2$ ,  $\Sigma_{\Theta}$  and  $\sigma_{rl}^2$
- Specify initial values for Θ<sub>c0</sub> and r<sub>lc0</sub>

Estimate full beliefs:

- Generate multiple realisations of model evolution
- Calculate empirical estimates for any expectations and (co-)variance structures of interest

- ▶ In particular:  $E(\mathbf{Y})$ ,  $Var(\mathbf{Y})$ ,  $Cov(\mathbf{Y}, \mathbf{\Theta})$
- ► Also:  $E(\Theta)$ ,  $Var(\Theta)$  ...

## The Bayes linear approach

Full Bayesian modelling of large systems:

- Difficult or impractical to make full prior specifications
- Non-physical simplifications required for modelling

Bayes linear modelling:

- Requires specification of partial beliefs only
- Is computationally efficient for high dimensional problems
- Uses expectation as a primitive rather than probability
- Beliefs are updated using adjusted expectations
- de Finetti [1974] or Goldstein and Wooff [2007]

# Adjusting beliefs

Observe data D to update beliefs B

The adjusted expectation vector for B given D is:

$$E_D(B) = E(B) + \operatorname{Cov}(B, D)\operatorname{Var}(D)^{\dagger}(D - E(D))$$

The **adjusted variance** matrix for *B* given *D* is:

 $\operatorname{Var}_D(B) = \operatorname{Var}(B) - \operatorname{Cov}(B, D)\operatorname{Var}(D)^{\dagger}\operatorname{Cov}(D, B)$ 

- ► *E<sub>D</sub>*(*B*) used as an **updated estimator** for *B*
- ► Var<sub>D</sub>(B) can be viewed as the mean square error of the estimator E<sub>D</sub>(B)

└─ The Bayes linear approach

## Motivating Bayes linear

Two collections of random quantities,  $B = (B_1 \dots B_r)$  and  $D = (D_1 \dots D_s)$ . The **adjusted expectation** for  $B_i$  given D is the linear combination  $a_i^T D$ ,

$$E_D(B) = \sum_{i=0}^s a_i^{\mathsf{T}} D_i$$

which minimises;

$$E\left((B_i-\sum_{i=0}^k a_i^T D_i)^2\right)$$

over choices of  $a_i^{T}$ .

Must specify prior mean vectors and variance matrices for B and D and a covariance matrix between B and D.

# Exchangeable events

- In an exchangeable sequence of random variables, future samples behave like earlier ones
- ► A collection of quantities X = {X<sub>1</sub>, X<sub>2</sub>,...} is exchangeable if our beliefs are invariant under permutation of X
- The role of exchangeability in subjective analysis is analogous to that of independence in classical inference
- An exchangeable sequence can be represented as a mixture of underlying i.i.d. sequences (de Finetti [1974])

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Updating beliefs
Exchangeable events

### Exchangeability and independence

Independent events are exchangeable, but exchangeable events may not be independent

- A sequence of i.i.d. random variables is exchangeable
- Sampling without replacement is exchangeable, but not independent
- ▶ For the bivariate normal random variable:

$$Z \sim N\left( \left(\begin{array}{c} 0\\ 0 \end{array}\right), \left(\begin{array}{c} 1 & \rho\\ \rho & 1 \end{array}\right) \right)$$

components  $Z_1$  and  $Z_2$  are exchangeable, but independent only if  $\rho = 0$ 

### Second order exchangeability

A collection  $X = \{X_1, X_2, ...\}$  is second order exchangeable if our beliefs about first and second order specification are invariant under permutation of X

$$E(X_i) = \mu$$
  $\operatorname{Var}(X_i) = \sigma$   $\operatorname{Cov}(X_i, X_j) = \gamma$   $i \neq j$ 

Equivalent to full exchangeability for Bayes linear modelling

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Updating beliefs
Exchangeable events

### The representation theorem

For (s.o.) exchangeable  $X = X_1, X_2, ...$ , we **represent** each  $X_i$  as the sum of two random quantities, a "**mean**" plus "**residual**":

$$X_i = \mathcal{M} + \mathcal{R}_i$$

Each pair  $\mathcal{R}_i$  and  $\mathcal{R}_j$  are **uncorrelated**  $i \neq j$  and each  $\mathcal{R}_i$  is uncorrelated with  $\mathcal{M}$  (Goldstein [1986])

$$\begin{aligned} E(\mathcal{M}) &= \mu & & \operatorname{Var}(\mathcal{M}) &= \gamma \\ E(\mathcal{R}_i) &= 0 & & \operatorname{Var}(\mathcal{R}_i) &= \sigma - \gamma \end{aligned}$$

- Simplifies specification of (co-)variance structures
- Adjust beliefs about  $\mathcal{M}$  not  $X_i$

# Exchangeable errors: simple (co-)variance structures

Global Effects: 
$$\Theta_{ct} = G\Theta_{ct-1} + \epsilon_{\Theta ct}$$
  $Var(\epsilon_{\Theta t}) = \Sigma_{\Theta}$ 

Assume (s.o.) exchangeability of  $\epsilon_{\Theta ct}$  over c and t

$$\epsilon_{\Theta ct} = \mathcal{M}_{\Theta} + \mathcal{R}_{\Theta ct}$$

- Then  $Var(\epsilon_{\Theta ct}) = \sigma_{\Theta}^2$ , for all c and t
- And  $\operatorname{Cov}(\epsilon_{\Theta c't'}, \epsilon_{\Theta ct}) = \gamma_{\Theta}$ , for all  $c' \neq c$  and  $t' \neq t$
- Hence, a simple **two parameter form** for  $\Sigma_{\Theta} = \Sigma_{\Theta}(\sigma_{\Theta}^2, \gamma_{\Theta})$

## Exchangeable squared errors: (co-)variance learning

Global Effects:  $\Theta_{ct} = G\Theta_{ct-1} + \epsilon_{\Theta ct}$   $Var(\epsilon_{\Theta t}) = \Sigma_{\Theta}$ Assume (s.o.) exchangeability of  $\epsilon_{\Theta ct}^2$  over c and t

$$\epsilon_{\Theta ct}^2 = \mathcal{M}_V + \mathcal{R}_{Vct}$$

- Then  $E(\epsilon_{\Theta ct}^2) = E(\mathcal{M}_V) = \sigma_{\Theta}^2$ , for all c and t
- Hence adjusting beliefs about M<sub>V</sub> allows us to learn about variances

- Updating beliefs

Exchangeable events

### Method components revisited

- 1. Specify model
  - Partial belief structure
  - Exchangeability assumptions (if any)
- 2. Simulate to estimate full belief structure
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  - Incomplete and irregular observations
  - ► Learn about system level and (co-)variance structure

- 4. Simulate adjusted system to forecast
- 5. Make decision
  - Expected loss to optimise decision

# Making decisions: Optimal inspection design

- Identify good inspection designs with which to update our beliefs
- Potential designs evaluated in terms of reducing uncertainty about critical system characteristics
- Utility or loss is used to compare designs

For example:

Simple decision to replace or retain a system component subject to potential costly failure

### Loss for component replacement

- Simple maintenance decision  $\delta \in \Delta$  to replace R or retain  $\overline{R}$ .
- Outcome  $o \in O$  is either failure F or survival  $\overline{F}$ .
- Loss  $L(o, \delta)$  is specified as:

$$\begin{array}{c|c} F & \overline{F} \\ \hline R & L_R & L_R \\ \overline{R} & L_F & 0 \end{array}$$

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Updating beliefs
Making decisions

### Expected loss with observed data

For **observed** data **D**:

 $E_{[O|D]}[L(O,\delta)|D] = L(F,\delta)\operatorname{Pr}(F|D) + L(\bar{F},\delta)\operatorname{Pr}(\bar{F}|D)$ 

$$\begin{split} E[L(O,R)|D] &= L(F,R)\mathrm{Pr}(F|D) + L(\bar{F},R)\mathrm{Pr}(\bar{F}|D) = L_R\\ E[L(O,\bar{R})|D] &= L(F,\bar{R})\mathrm{Pr}(F|D) + L(\bar{F},\bar{R})\mathrm{Pr}(\bar{F}|D) = L_F\mathrm{Pr}(F|D) \end{split}$$

Replacement is selected when:

$$E[L(O, R)|D] < E[L(O, \overline{R})|D]$$
$$\Pr(F|D) > \frac{L_R}{L_F}$$

### Expected loss with unobserved data

Expected loss of decision  $\delta$  based on **as yet unobserved** data *D* from **design** *d* is:

$$\begin{split} E_{[O]}[L(O,\delta)] &= E_{[D]}\{E_{[O|D]}[L(O,\delta)|D]\}\\ &= E_{[D]}\{L(F,\delta)\mathrm{Pr}(F|D) + L(\bar{F},\delta)\mathrm{Pr}(\bar{F}|D)\} \end{split}$$

**Optimal decision**  $\delta^*$  satisfies:

$$\delta^* = \begin{cases} R \text{ if } \Pr(F|D) > \rho \\ \bar{R} \text{ if } \Pr(F|D) \le \rho \end{cases} \quad \text{where} \quad \rho = \frac{L_R}{L_F}$$

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 $= L_R I_1 + L_F I_2$ 

- +  $L_F E\{\Pr(\Pr(F|D) | \Pr(\Pr(F|D) \le \rho)\}\Pr(\Pr(F|D) \le \rho)\}$
- $= L_R \Pr(\Pr(F|D) > \rho)$
- +  $E\{L(F, \delta^*) \Pr(F|D) + L(\overline{F}, \delta^*) \Pr(\overline{F}|D) | \delta^* = \overline{R}\} \Pr(\delta^* = \overline{R})$

- $= E\{L(F,\delta^*)\Pr(F|D) + L(\bar{F},\delta^*)\Pr(\bar{F}|D)|\delta^* = R\}\Pr(\delta^* = R)$
- $= E_{[D]}\{L(F,\delta^*)\Pr(F|D) + L(\bar{F},\delta^*)\Pr(\bar{F}|D)\}$
- $= E_{[D]} \{ E_{[O|D]} [L(O, \delta^*) | D] \}$

Expected loss for **design**,  $E[L(O, \delta^*)]$  $E_{[O]}[L(O, \delta^*)]$ 

Updating beliefs

Learning about large industrial systems

# Expected loss for **design**, $E[L(O, \delta^*)]$

$$E[L(O, \delta^*)] = L_R I_1 + L_F I_2$$

- Integrals *I*<sub>1</sub> and *I*<sub>2</sub> evaluated for given probability distributions characterised by location and scale parameters
- Adjusted expectations and variances from the Bayes linear update used to estimate location and scale
- Computationally fast: no need to simulate data D for given design d

Learning about large industrial systems
Application: corrosion monitoring
Corrosion monitoring

# Application: Corrosion monitoring of offshore platform



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Application: corrosion monitoring
Corrosion monitoring

# Corrosion monitoring

- Offshore platforms have large numbers of components subject to corrosion
- Corrosion can lead to failure incurring costs
- A typical offshore platform has >100 corrosion circuits, each with 20 to 1000 components, hence potentially >5000 components subject to corrosion.

Some corrosion circuits have similar characteristics

Learning about large industrial systems — Application: corrosion monitoring

Corrosion monitoring

### Typical corrosion circuit diagram



### Data characteristics

- Minima: over whole component observed
- Short time series: data per component is limited, but large number of components
- Irregular inspections: inspections are carried out when possible, often when processes are shut down, often several months or years apart
- Incomplete inspections: due to size of systems and inaccessibility of components, complete systems are rarely inspected

### Typical inspection design for a corrosion circuit



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Application: corrosion monitoring

Data characteristics

### Method components

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### Model

The system is modelled as:

$$\begin{aligned} Y_{tc} &= \min_{l} \left( X_{tc} + r_{tcl} + \epsilon_{Ytcl} \right) & \operatorname{Var}(\epsilon_{Ytcl}) = \sigma_{Yc}^{2} \\ X_{tc} &= X_{t-1c} + \alpha_{tc} + \epsilon_{Xtc} & \operatorname{Var}(\epsilon_{Xtc}) = \Sigma_{X} \\ \alpha_{tc} &= \alpha_{t-1c} + \epsilon_{\alpha tc} & \operatorname{Var}(\epsilon_{\alpha tc}) = \Sigma_{\alpha} \\ r_{tcl} &= r_{t-1cl} + \epsilon_{rtcl} & \operatorname{Var}(\epsilon_{rtcl}) = \sigma_{rc}^{2} \end{aligned}$$

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# Learning about wall thickness and corrosion rate

- Perform simulations of model based on partial belief specification
- Simulations together with inspection data yield updated adjusted expectations for wall thickness and corrosion rate parameters

 Modelling covariance structure, we learn about all components even unobserved

### Typical covariance structure based on adjacency



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# Variance learning: why?

- Prior specification of (co-)variances is difficult
- Variance parameters in model typically fixed. Poor prior specification leads to poor model performance
- Variance is not directly observable. Adjusting beliefs more difficult

### Variance learning: simple corrosion model For example:

$$X_{ct} = X_{ct-1} + \alpha_{ct} + \epsilon_{Xct}$$
$$\alpha_{ct} = \alpha_{ct-1} + \epsilon_{\alpha ct}$$

**Differences of observations** eliminate effects of wall thickness' and corrosion rates (Wilkinson [1997])

$$X_t^{(1)} = X_{ct} - X_{ct-1} = \alpha_{ct} + \epsilon_{Xct} = \alpha_{ct-1} + \epsilon_{\alpha ct} + \epsilon_{Xct}$$
$$X_t^{(2)} = X_{ct} - X_{ct-2} = X_{ct-1} + \alpha_{ct} - X_{ct-2} + \epsilon_{Xct}$$
$$= \alpha_{ct} + \alpha_{ct-1} + \epsilon_{Xct} + \epsilon_{Xct-1}$$
$$= 2\alpha_{ct-1} + \epsilon_{\alpha ct} + \epsilon_{Xct} + \epsilon_{Xct-1}$$

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### Variance learning: squared differences

Therefore:

$$X_t^{(2)} - 2X_t^{(1)} = -\epsilon_{\alpha ct} - \epsilon_{Xct} + \epsilon_{Xct-1}$$

and:

$$E[(X_t^{(2)} - 2X_t^{(1)})^2] = E[(-\epsilon_{\alpha ct} - \epsilon_{Xct} + \epsilon_{Xct-1})^2]$$
$$= E[\epsilon_{\alpha ct}^2] + E[\epsilon_{Xct}^2] + E[\epsilon_{Xct-1}^2]$$
$$= \sigma_{\alpha c}^2 + 2\sigma_{Xc}^2$$

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### Variance learning: exchangeability in time

Assume squares of residuals are (s.o.) exchangeable in time. Using representation theorem:

$$[\epsilon_{Xct}]^2 = \mathcal{M}(V_c) + \mathcal{R}_t(V_c)$$

where:

$$E([\epsilon_{Xct}]^2) = \sigma_{Xc}^2 = V_c \quad \text{Var}([\epsilon_{Xct}]^2) = \Sigma_{V_c}$$
$$Cov([\epsilon_{Xct}]^2, [\epsilon_{Xct'}]^2) = \Gamma_{V_c} \quad t \neq t'$$

## Variance learning: adjusting beliefs

Compute  $E_D[\mathcal{M}(V_c)]$ :

$$D = \left\{ \frac{(X_t^{(2)} - 2X_t^{(1)})^2}{2 + \lambda} \right\}_{t=3}^T$$

 $E_D[\mathcal{M}(V_c)] = E[\mathcal{M}(V_c)] + \operatorname{Cov}[\mathcal{M}(V_c), D]\operatorname{Var}[D]^{-1}(D - E(D))$  $= \sigma_{Xc}^2 + 2'_T \Gamma_{V_c} \operatorname{Var}[D]^{-1}(D - 1_T(\sigma_{\alpha c}^2 + 2\sigma_{Xc}^2))$ 

yielding an adjusted estimate for the variances in the model

Application: corrosion monitoring

Bayes linear variance learning

# Variance learning: generalisations

Generalisations include:

- General time step form for irregular time points
- Partial inspections using exchangeable variances across components
- Mahalanobis distance fitting to update local variances

Learning about large industrial systems
Application: corrosion monitoring
Model diagnostics

# Model diagnostics

 Mahalanobis distance to estimate data discrepancy, comparing data to our prior estimates

$$\mathrm{Dis}(X) = \frac{(D - E(D))^2}{\mathrm{Var}D}$$

 For each of our updated values we can also compute the adjustment discrepancy

$$\operatorname{Dis}_{D}(X) = \frac{(E_{D}(X) - E(X))^{2}}{\operatorname{RVar}_{D}X}$$

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Learning about large industrial systems Application: corrosion monitoring

Model diagnostics

# Typical model diagnostics



Learning about large industrial systems — Conclusions and future work

### Conclusions

General purpose framework for modelling and inspection design of large systems

Compared to existing methods, the model is novel in that:

- Analysis of multivariate systems possible, rather than modelling components separately and independently
- Data from incomplete inspections at arbitrary times used to learn about the whole system
- Uncertainties in system parameters adjusted, as are the dependencies between these
- Economically-optimal future inspection strategies can be estimated consistently

Learning about large industrial systems — Conclusions and future work

### Future work

- Efficient implementation of sequential Bayes linear calculation
- Search methods for good designs in high dimensions
- Elicitation of prior partial beliefs
- Flexible forms for modelling for system element behaviour
- Enhanced criteria for evaluation of inspection schemes
- Fundamental modelling of physical processes (e.g. corrosion)
- New applications to manufacturing, environmental and commercial problems

### Thank you

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### Backup

$$E_{[Y]}(g(Y)) = E_{[X]}(E_{[Y|X]}(g(Y)|X))$$
$$E_{[Y|X]}(g(Y)|X) = \sum_{i} g(Y_i) \operatorname{Pr}(Y_i|X)$$

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