Recent Progress on Bayesian Decision-Theoretic Clinical Trial Designs

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Clinical Trials

- Two treatments: control (existing one) and novel (not approved yet)
- Is the novel treatment better than the control?
 - clinically relevant treatment effect difference
 if not, it will not be approved!
- One sets up a clinical trial of n patients
- A (clinical trial) design is an allocation policy that specifies which treatment the *i*th patient will receive

Clinical Trials

The gold standard design: randomised controlled trial
 50% vs 50% fixed equal randomisation
 avoids all types of biases
 in use since 1948 (advocated since Hill 1937)

- Its main goal is to learn about treatment effectiveness with a view to prioritising future outside patients
 - maximises power of a treatment effect difference
 if approved, future patients are, say, 95% confident that the novel treatment is better than the control
- A half of trial patients gets the inferior treatment

Randomised Controlled Trial

Advantages of randomised controlled trial

- best possible for estimating the treatment effect
- "easy" to understand by trial statisticians, physicians
 - basic statistical knowledge
- ▷ "quick" to design
 - trial size formulae available
- "straightforward" to implement
 - no computation needed during the trial
- ▷ "easy" to understand by in-trial patients (?)
- ▷ "easy" to interpret the results by regulators
- ▷ "easy" to interpret the results by physicians (?)
- ▷ "easy" to interpret the results by patients (?)

Randomised Controlled Trial

- Disadvantages of randomised controlled trial
 - \triangleright cost: 20% error of not approving a better treatment
 - development and approval processes: \$ billions
 - Faith: once approved, no (simple) way to retract the treatment
 - worse treatment approved by 5% chance
 - unforeseen long-term secondary effects
 - Feasibility: requires hundreds of patients for a trial
 - stationarity: approval process takes years
 - inappropriate for new diseases and epidemics
 - ethics: patients join a trial expecting to get a possibly better (unapproved yet) treatment

Randomised Controlled Trial

- Statistical testing based on randomised equal allocation is a widespread state-of-the-art approach in the design of experiments, under different names:
 - randomised controlled trial in biostatistics
 between-group design in social sciences
 A/B testing in Internet marketing
- In other fields however the disadvantages are usually less severe

Bayesian Decision-Theoretic Trial

"...there can be no objection to the use of data, however meagre, as a guide to action required before more can be collected ... Indeed, the fact that such objection can never be eliminated entirely—no matter how great the number of observations—suggested the possible value of seeking other modes of operation than that of taking a large number of observations before analysis or any attempt to direct our course... This would be important in cases where either the rate of accumulation of data is slow or the individuals treated are valuable, or both."

Bayesian Decision-Theoretic Trial

- Proposed in Thompson 1933 (pre-dates Hill 1937)
- The goal is to provide higher health benefit to both in-trial patients and after-trial patients
 - as opposed to the RCT's learning goal of reliable treatment effect estimation
- It is done by deciding the allocation, i.e., the randomisation probabilities for every patient (or for a group of patients)
 - response-adaptive: decisions based on the responses accumulated so far, i.e. Bayesian

Bayesian Decision-Theoretic Trial

- In theory, can be solved to optimality by decision theory
- In practice, optimal decisions are computed numerically
 - it is often believed tractable only for small trials
- Milestones IMHO
 - Thompson (Biometrika 1933)
 - ▷ Glazebrook (Biometrika 1978)
 - ▷ Gittins & Jones (Biometrika 1979)
 - ▷ Armitage (ISR 1985)
 - Cheng, Su & Berry (Biometrika 2003)
 - Berry (Nature 2006), Cheng & Berry (Biometrika 2007)
 - ▷ Villar, Bowden & Wason (Statistical Science 2015)

Health Benefit Approach

- Important because healing patients is the ultimate goal of new treatment development
- Bayesian decision-theoretic model
 - optimally solving learning/healing trade-off
 both learning and healing takes place during the trial
- This kind of general problem became known as the multi-armed bandit problem

Multi-Armed Bandit Problem



Multi-armed Bandit Problem

- Optimally solving learning/earning trade-off
- Studied in other scientific disciplines including probability, statistics, operational research, economics, marketing, machine learning, computer simulation, computer science, and communications engineering
- Many formulations: i.i.d., Markovian, Bayesian, etc.
- Many extensions, mainly in machine learning
- Appropriate model for trials: finite horizon
 - after the end of the horizon we will not be able to influence the allocation

Bayesian Bernoulli Bandit Model

- Finite horizon: n sequentially arriving patients
- Two-armed: treatment A or B for each patient
- Binary endpoints: success (1) or failure (0)
- Let X_i and Y_i denote patient i's response from treatment A and B respectively (for i = 1,...,n). Then,

 $X_i \sim \text{Bernoulli}(1, \theta_A)$ and $Y_i \sim \text{Bernoulli}(1, \theta_B)$,

where θ_A and θ_B are the unknown success probabilities of treatments A and B respectively

Bayesian Approach

- Beliefs $\widehat{ heta}_A$ and $\widehat{ heta}_B$ to be updated over the trial
- Prior Distribution: $\hat{\theta}_A \sim \text{Beta}(a, b)$, $\hat{\theta}_B \sim \text{Beta}(c, d)$ where we take a = b = c = d = 1 (uninformative)
- Posterior Distribution: After observing i (j) successes (failures) on treatment A, and k (l) successes (failures) on treatment B, the posterior distribution is represented by another Beta distribution (by conjugacy)

 $\widehat{\theta}_A \sim \mathsf{Beta}(a+i,b+j)$, $\widehat{\theta}_B \sim \mathsf{Beta}(c+k,d+l)$

DP Design

- We use dynamic programming (DP) to obtain an optimal adaptive treatment allocation sequence
- Optimal in the sense of maximising the expected total number of successes in the trial
- Specifically, we use backward induction algorithm
- Let $\mathcal{F}_m(i, j, k, l)$ be the expected total number of successes under an optimal policy after m patients
- If m = n, there is nothing to do: $\mathcal{F}_n(i, j, k, l) = 0$ $\forall i, j, k, l$

Backward Induction

- If m = n 1 (one patient left):
 - 1. If treatment A, we compute the expectation

$$\mathcal{F}_{n-1}^A(i,j,k,l) = \frac{i}{i+j} \cdot 1 + \frac{j}{i+j} \cdot 0$$

2. If treatment B, we compute the expectation

$$\mathcal{F}_{n-1}^B(i,j,k,l) = \frac{k}{k+l} \cdot 1 + \frac{l}{k+l} \cdot 0$$

We wish to choose the optimal allocation such that

$$\mathcal{F}_{n-1}(i, j, k, l) = \max\{\mathcal{F}_{n-1}^{A}(i, j, k, l), \mathcal{F}_{n-1}^{B}(i, j, k, l)\}$$

DP Design

- Problem? This design is not suitable to implement in practice because it is completely deterministic
- As a result, there is a risk of introducing bias into the trial through the intentional selection of patients (selection bias)
- Therefore, we modify the DP design by forcing actions to be randomised
 - ▷ see also Cheng & Berry (Biometrika, 2007)
- Helps to maintain blinding and reduce the risk of bias

Forcing Randomised Actions

- Action 1: treatment A is allocated with probability p
- Action 2: treatment B is allocated with probability p
- The expected total number of successes under Action 1

$$\mathcal{V}_m^1(i,j,k,l) = p \cdot \mathcal{F}_m^A(i,j,k,l) + (1-p) \cdot \mathcal{F}_m^B(i,j,k,l)$$

• The objective function becomes

 $\mathcal{V}_m(i,j,k,l) = \max\left\{\mathcal{V}_m^1(i,j,k,l), \ \mathcal{V}_m^2(i,j,k,l)\right\}$

• Lower selection bias, but lower controllability

Randomised Variant

• Problems? After running simulations, we found:

- this design is very underpowered for high p
 in some of the runs (only a few out of 10,000), all patients were allocated to only one of the treatments
- This means we cannot be confident about the results
- ...we cannot calculate important performance measures
- Therefore, we lower-limit the number of observations on each treatment

Constrained Variant

- We modify the optimal randomised DP policy by adding a constraint to ensure that we obtain ≥ ℓ observations from each treatment
- To do this, we assign a large penalty to every terminal state that has $< \ell$ observations on a treatment arm
- The undesirable states will now be avoided (as much as possible) by the optimal policy
- We tried a range of values for ℓ, i.e. 0.05n, 0.10n, 0.15n, 0.20n and 0.25n. (Note that 0.50n corresponds to equal, fixed randomisation)

Simulation Study

• We evaluate the performance of proposed designs by

- Bias of the treatment effect estimator
- ▷ ...and its mean squared error (MSE)
- Statistical power
- Expected proportion of successes (EPS)
- Patients allocated to the superior arm (On sup)
- For each configuration, we replicate 10,000 trials

Simulation Results: Randomised Variant

Example. $n = 75, \theta_A = 0.2, \theta_B = 0.8$

p	Bias	MSE	Power	EPS	On sup
50%	0.001	0.004	1.000	50.0%	50.0%
60%	0.001	0.005	1.000	55.7%	59.6%
70%	0.001	0.007	0.999	61.5%	69.2%
80%	0.004	0.010	0.995	67.2%	78.8%
90%	0.009	0.019	0.937	73.0%	88.3%
100%	0.100	0.043	0.118	78.6%	97.6%

• The Power (almost) does not change if p increased from 50% to 60% or 70%. Room for increasing EPS!

Simulation Results: Randomised Variant

Example. $n = 75, \theta_A = 0.2, \theta_B = 0.6$

p	Bias	MSE	Power	EPS	On sup
50%	0.001	0.004	0.938	40.0%	50.0%
60%	0.002	0.005	0.935	43.7%	59.1%
70%	0.002	0.007	0.910	47.3%	68.2%
80%	0.005	0.009	0.830	50.9%	77.3%
90%	0.015	0.015	0.636	54.4%	86.0%
100%	0.089	0.030	0.070	57.7%	94.2%

• The Power (almost) does not change if p increased from 50% to 60% or 70%. Room for increasing EPS!

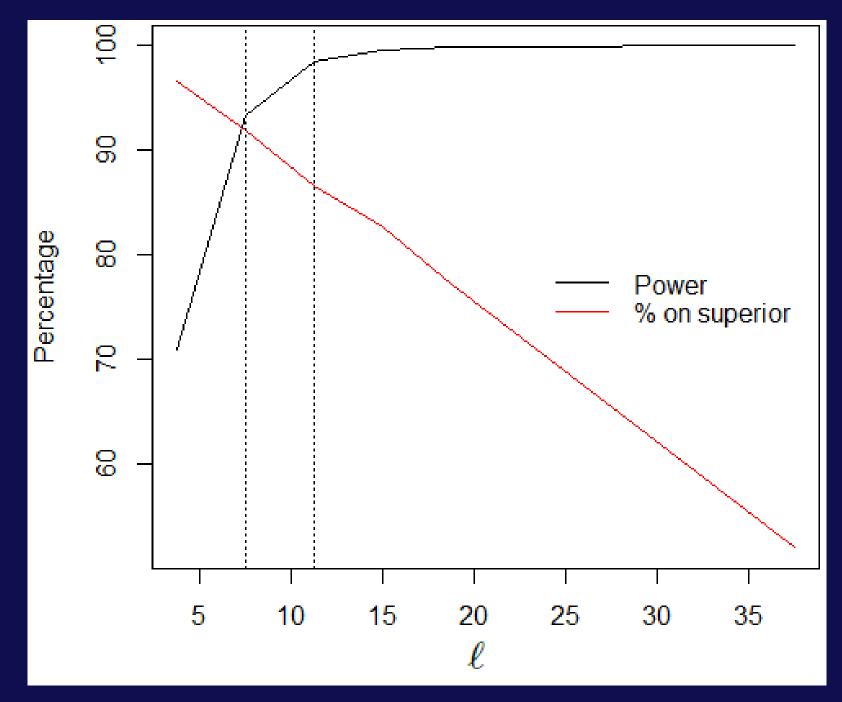
Simulation Results: Constrained Variant

Example. $n = 75, \theta_A = 0.2, \theta_B = 0.8$

ℓ	Power	EPS	On sup
0.05n	0.442	78.0%	96.6%
0.10n	0.884	75.2%	91.9%
0.15n	0.964	72.1%	86.7%
0.20n	0.985	69.7%	82.7%
0.25n	0.997	66.5%	77.3%
0.50n	1.000	51.2%	52.0%

 As l increases, the power of the design increases hyperbolically, but the EPS and % allocated to the superior arm decreases linearly

Simulation Results: Constrained Variant



Simulation Results: Constrained Randomised Variant

Example. $n = 75, \ell = 0.15n, \theta_A = 0.2, \theta_B = 0.8$

p	Bias	MSE	Power	EPS	On sup
60%	0.001	0.005	1.000	55.7%	59.6%
70%	0.001	0.007	0.999	61.5%	69.2%
80%	0.003	0.010	0.996	67.2%	78.7%
90%	0.003	0.014	0.977	71.3%	85.5%

• The Power is quite high even if p increased to 80% or 90%. Also bias diminishes!

Simulation Results: Constrained Randomised Variant

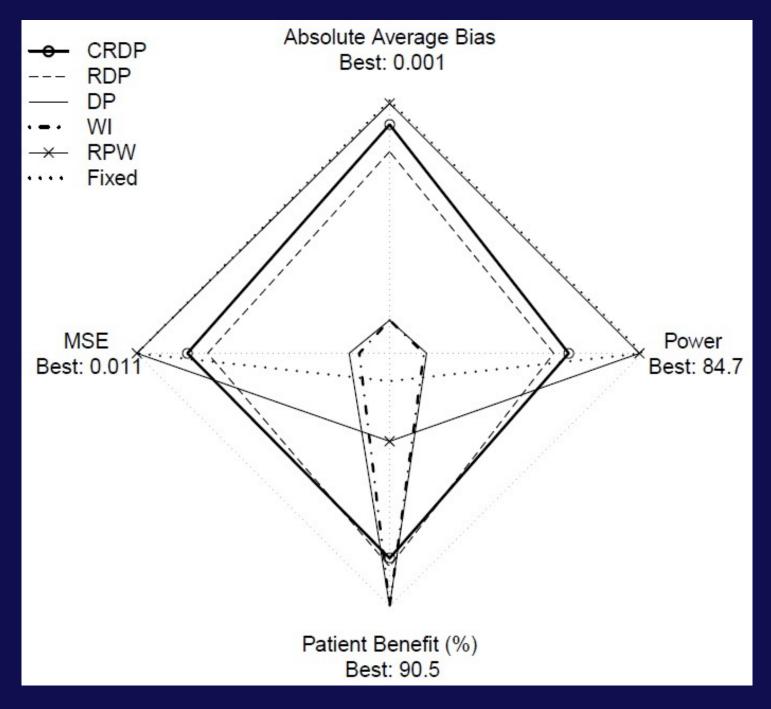
Example. $n = 75, \ell = 0.15n, \theta_A = 0.2, \theta_B = 0.6$

p	Bias	MSE	Power	EPS	On sup
60%	0.002	0.005	0.935	43.7%	59.1%
70%	0.002	0.007	0.910	47.3%	68.2%
80%	0.005	0.009	0.834	50.9%	77.2%
90%	0.008	0.013	0.724	53.6%	84.0%

• The Power is quite high even if p increased to 80% or 90%. Also bias diminishes!

Simulation Study

- We compare our proposed constrained randomised variant of DP (CRDP) design to the following designs:
 - Fixed randomisation (the gold standard)
 - Randomised play-the-winner rule (RPW)
 - Optimal dynamic programming policy (DP)
 - ▷ Whittle index policy (WI)
 - Randomised variant of the DP policy (RDP)
- We suggest to set $p = 90\%, \ell = 0.15n$ in CRDP



Example. $n = 75, \theta_A = 0.2, \theta_B = 0.8$

Design	Bias	Power	EPS	On sup
Fixed	0.000	1.000	50.0%	50.0%
RPW	0.008	0.998	66.2%	76.9%
WI	0.098	0.108	78.6%	97.6%
DP	0.100	0.118	78.6%	97.5%
RDP	0.009	0.937	73.0%	88.3%
CRDP	0.003	0.977	71.3%	85.5%

Example. $n = 75, \theta_A = 0.2, \theta_B = 0.6$

Design	Bias	Power	EPS	On sup
Fixed	0.000	0.935	40.0%	50.0%
RPW	0.002	0.928	46.2%	65.4%
WI	0.092	0.066	57.8%	94.4%
DP	0.088	0.074	57.7%	94.1%
RDP	0.015	0.636	54.4%	86.0%
CRDP	800.0	0.724	53.6%	84.0%

- In our suggested CRDP design
 - b the % expected proportion of successes is much higher than in the traditional fixed and RPW designs
 - b the % allocated to the superior arm is much higher than in the traditional fixed and RPW designs
 - b the power is largely improved upon relative to the other bandit designs
 - the bias is negligible, opposed to large bias of other bandit designs

Conclusion about CRDP

- We address some of the key issues preventing bandit models from being implemented in clinical trial practice
 - lack of randomisation
 - insufficient statistical power
 - biased estimates of the treatment effect
- We need to talk to statisticians and clinicians about bandit models
 - give me randomisation probability and desired power
 I tell how to randomise treatments to heal patients
- See Williamson et al. (CSDA 2017)

Optimal Designs

- The designs computed using DP (DP, RDP, CRDP) are optimal, i.e. provide the maximum health benefit given their respective restrictions
- For the two-armed case (on a standard laptop):
 - \triangleright a basic R code can design trials of size up to 200
 - \triangleright an efficient Julia code up to 1,000
 - a during-the-trial computation allows even larger trials
- Longer trials can be designed on a workstation/cloud
- More complex trials with much smaller sizes

Near-optimal Designs

- Near-optimal designs can be obtained using index rules
 - idea: for every patient, allocate the treatment which has the highest value of an adaptive priority index
 - these indices are obtained in a particular way (Gittins index, Whittle index), and are asymptotically optimal
 - suboptimality is negligible
 - ▷ do not suffer from the curse of dimensionality
 - suitable for more complex trials
- Unfortunately, not enough time to explain in detail
 - requires understanding of Markov decision processes

Near-optimal Designs

• See e.g.,

- Glazebrook (Biometrika 1978)
- ▷ Gittins & Jones (Biometrika 1979)
- Coad (Biometrika 1991)
- ▷ Wang (Biometrika 1991)
- ▷ Villar, Bowden & Wason (Statistical Science 2015): survey
- ▷ Villar, Wason & Bowden (Biometrics 2015): group-sequential
- ▷ Villar (PEIS 2017): Whittle index rule
- Villar & Rosenberger (Biometrics 2017): covariates
- ▷ Villar, Bowden & Wason (Pharmaceutical S. 2018): time trends
- ▷ Smith & Villar (Applied S. 2018): Gaussian responses
- Williamson & Villar (2018+): Gaussian responses & group-sequential

Optimal and Near-optimal Designs

- Neither the optimal nor near-optimal designs have been implemented in practice yet
 - ▷ will Roche be the first?

Thompson's Posterior Sampling Designs

- Thompson 1933 proposed a heuristic:
 - randomise according to the posterior probability of being the best arm
- This can be done by exact calculation or by sampling
- Recently, several trials have been designed in this way
 - Don Berry (MD Anderson) and Berry Consultants
 e.g. I SPY-2, GBM Agile
- Several recent papers by a group at Harvard
- "Easy" to use, but quite suboptimal

Related Designs

- To overcome the computational difficulty of optimal designs and to add randomisation, improve estimation, etc., several designs have been proposed
 - ▷ not in Bayesian framework
- Urn designs (e.g. Randomised Play-the-Winner)
 - based on the property of the optimal design to allocate the treatment again if a success was observed on the last patient
 - there are several variants (e.g. Drop-the-Loser)
 well studied theoretically (e.g. Rosenberger)

Related Designs

- Biased-coin designs
 - well studied theoretically
- Relatively easy to use, but too suboptimal
 - \triangleright randomisation probabilities not too far from 50%

Conclusion about Designs

- Both the Thompson's posterior sampling and the above related designs are myopic
 - randomisation decisions are bases on the assumption that the next patient is the last one
- Optimal designs and index rules are not myopic
 - randomisation decisions take into account the remaining trial size and the after-trial population size
 this is the reason why they provide a significantly higher health benefit

Work in Progress

- With F. Williamson and T. Jaki, we are studying the CRDP design if there are delayed responses
- With J. Wason, we are looking at when it is optimal to add in a novel treatment to a platform trial
- I am working on a Julia package to compute the optimal designs

Thank you for your attention