

STOR608 Sprint 1: Optimisation Under Uncertainty

Group 2:

Theo Crookes, Max Howell, Robert Lambert, James Neill
Sprint Lead: Jamie Fairbrother

25 November 2022

Network Planning with Random Demand

- We construct a model for networks that provide private line services.

Network Planning with Random Demand

- We construct a model for networks that provide private line services.
- Due to the stochastic nature of the demands, they are treated as random variables.

Network Planning with Random Demand

- We construct a model for networks that provide private line services.
- Due to the stochastic nature of the demands, they are treated as random variables.
- In order to carry out an optimisation, we treat the problem as a two-stage stochastic programme.

Network Planning with Random Demand

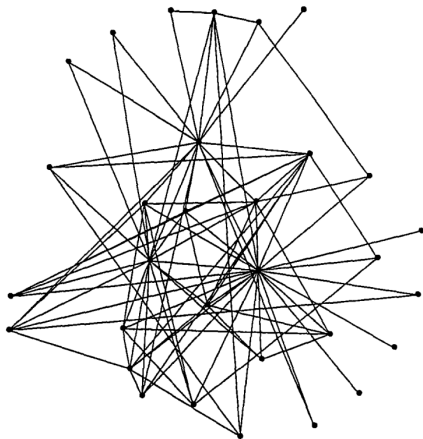


Figure: Graphical representation of the network planning problem Sen et al. (1994).

Notation

Symbol	Explanation
n	Number of links/edges in the graph
$x \in \mathbb{R}^n$	A vector denoting the amount of additional capacity for each link j
m	The number of point-to-point pairs served by the network
$\tilde{d} \in \mathbb{R}^m$	A r.v. that represents the demands for the m point-to-point pairs
b	Total capacity in the system

More Notation

Symbol	Explanation
$R(i)$	denote a set of routes (r) that can be used for connections associated with point-to-point pair i
$A_{ir} \in \mathbb{R}^n$	An indicator vector where $(A_{ir})_j$ is 1 if link j is in route r (0 otherwise)
$e \in \mathbb{R}^n$	The current link capacities for each link j
f_{ir}	Number of connections associated with pair i using route $r \in R(i)$
s_i	Number of unserved requests for each pair i

Model Summary (Level 1)

Using this notation, we formulate the initial problem in the following way:

$$\min_x \mathbb{E}[h(x, \tilde{d})],$$

subject to

$$\sum_{j=1}^n x_j \leq b,$$
$$x \geq 0.$$

Model Summary (Level 2)

Within the first level of the model we must optimise the following linear programme:

$$h(x, d) = \min_{s, f} \sum_{i=1}^m s_i,$$

subject to

$$\sum_i \sum_{r \in R(i)} A_{ir} f_{ir} \leq x + e,$$

$$\sum_{r \in R(i)} f_{ir} + s_i = d_i,$$

$$f_{ir}, s_i \geq 0,$$

for all $i \in \{1, \dots, m\}$ and $r \in R(i)$.

Deterministic Equivalent Form

We can solve the previous problem as a single linear programme by converting the problem into Deterministic Equivalent form over T scenarios, each with probability $p^{(t)}$:

$$\min_{x,s,f} \left\{ \sum_{t=1}^T p^{(t)} \left(\sum_{i=1}^m s_i^{(t)} \right) \right\}$$

subject to

$$\begin{aligned} \sum_{i=1}^m \sum_{r \in R(i)} A_{ir} f_{ir}^{(t)} &\leq x + e, & \sum_{j=1}^n x_j &\leq b, \\ \sum_{r \in R(i)} f_{ir}^{(t)} + s_i^{(t)} &= d_i^{(t)}, & x, f, s &\geq 0, \end{aligned}$$

for all $i \in \{1, \dots, m\}$, $t \in \{1, \dots, T\}$.

Python Implementation

Using the previous formulation allows the problem to be minimised using a single optimisation over the three variables x, s, f .

We chose $p^{(t)} = 1/T$ as we have no prior information on the likelihood of individual scenarios.

Python Implementation

Using the previous formulation allows the problem to be minimised using a single optimisation over the three variables x, s, f .

We chose $p^{(t)} = 1/T$ as we have no prior information on the likelihood of individual scenarios.

Conditional on the number of scenarios T , the number of constraints is $T(n + m) + 1$ (excluding non-negativity).

Increasing the number of scenarios reduces the variability within the model, at the cost of increased time complexity. The increase in time taken will limit our computational capabilities going forward.

Time Taken

Here we see that as the number of scenarios increases, the time taken increases non-linearly.

Time Taken

Here we see that as the number of scenarios increases, the time taken increases non-linearly.

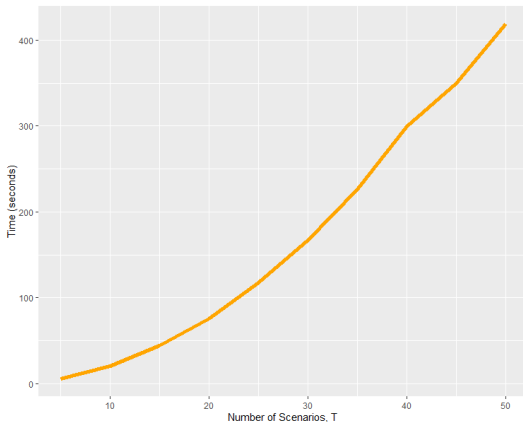


Figure: Optimisation times for varying T

Optimality Gap

We find some feasible value of x using a large value of T' ,

$$\tilde{x} = \left\{ x \in \mathcal{X} : \min_{x,s,f} \frac{1}{T'} \sum_{t=1}^{T'} h(x, \tilde{\xi}) \right\}.$$

Optimality Gap

We find some feasible value of x using a large value of T' ,

$$\tilde{x} = \left\{ x \in \mathcal{X} : \min_{x,s,f} \frac{1}{T'} \sum_{t=1}^{T'} h(x, \tilde{\xi}) \right\}.$$

Then optimality gap is defined as

$$G_T^i = \frac{1}{T} \sum_{t=1}^T h(\tilde{x}, \xi_{it}) - \min_{x,s,f} \frac{1}{T} \sum_{t=1}^T h(x, \xi_{it}).$$

Optimality Gap Results

For sufficiently large T , we know that over k macroreplications

$$\frac{\bar{G}_T - \mathbb{E}(G_T)}{\frac{\sigma_T}{\sqrt{k}}} \sim t_{k-1},$$

where

$$\bar{G}_T = \frac{1}{k} \sum_{i=1}^k G_T^k,$$

$$\sigma_T^2 = \frac{1}{k-1} \sum_{i=1}^k (G_T^i - \bar{G}_T)^2.$$

Optimality Gap Results

For sufficiently large T , we know that over k macroreplications

$$\frac{\bar{G}_T - \mathbb{E}(G_T)}{\frac{\sigma_T}{\sqrt{k}}} \sim t_{k-1},$$

where

$$\bar{G}_T = \frac{1}{k} \sum_{i=1}^k G_T^i,$$
$$\sigma_T^2 = \frac{1}{k-1} \sum_{i=1}^k (G_T^i - \bar{G}_T)^2.$$

This means we can calculate a one-sided confidence interval for $\mathbb{E}(G_T)$:

$$\left(0, \bar{G}_T + t_{k-1, \alpha} \frac{\sigma_T}{\sqrt{k}} \right).$$

Optimality Gap Results

Using $k = 50$ macroreplications and $T = 20$ scenarios we obtain estimates $\bar{G}_T = 5.010$ and $\sigma_T = 0.908$.

Optimality Gap Results

Using $k = 50$ macroreplications and $T = 20$ scenarios we obtain estimates $\bar{G}_T = 5.010$ and $\sigma_T = 0.908$.

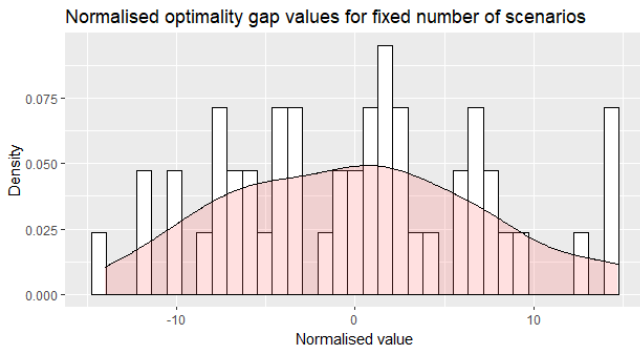
This means we have a 95% confidence interval for $\mathbb{E}(G_T)$ of $(0, 5.225)$.

Optimality Gap Results

Plotting a histogram of the standardised outputs, we expect the results to follow the distribution of a t -distribution on $k - 1 = 49$ degrees of freedom.

Optimality Gap Results

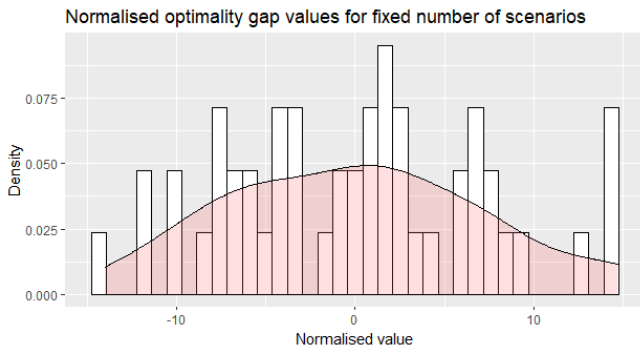
Plotting a histogram of the standardised outputs, we expect the results to follow the distribution of a t -distribution on $k - 1 = 49$ degrees of freedom.



Unfortunately this is not the case with our data.

Optimality Gap Results

Plotting a histogram of the standardised outputs, we expect the results to follow the distribution of a t -distribution on $k - 1 = 49$ degrees of freedom.



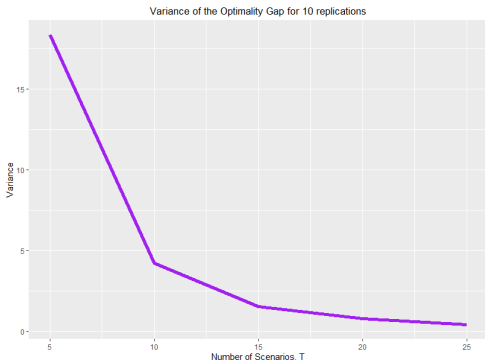
Unfortunately this is not the case with our data. :(

Optimality Gap Results

For fixed number of macroreplications, as T increases, the variance of the estimate of the optimality gap tends to 0.

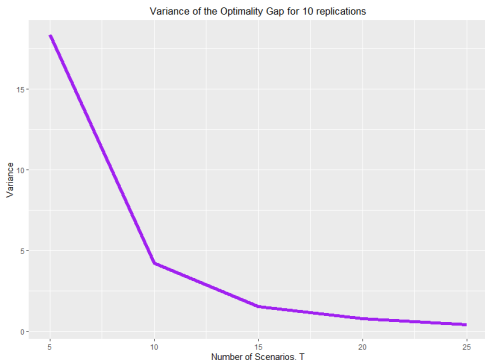
Optimality Gap Results

For fixed number of macroreplications, as T increases, the variance of the estimate of the optimality gap tends to 0.



Optimality Gap Results

For fixed number of macroreplications, as T increases, the variance of the estimate of the optimality gap tends to 0.



Therefore, as T increases, the estimate of the optimality gap tends to the true optimality gap.

Risk reduction methods

Possible methods to make the formulation more risk averse:

1. VaR.
2. CVaR
3. Minimise Variance.
4. Minimise Semivariance.
5. Minimise the squared losses.
6. More recent proposals e.g. most probable maximum size of risk events (MPMR) [Chen and Cheng (2022)]

β -CVaR

β -CVaR is defined as

$$\beta\text{-CVaR}(\theta) = \int_{\beta}^1 F^{-1}(u) du$$

which can be thought of as the expected loss if the worst case threshold is every passed.

β -CVaR

β -CVaR quantifies the amount of tail risk in the model.

β -CVaR

β -CVaR quantifies the amount of tail risk in the model.

β -CVaR is tractable.

β -CVaR

β -CVaR quantifies the amount of tail risk in the model.

β -CVaR is tractable.

From Acerbi and Tasche (2002) we know that β -CVaR is a coherent risk measure, whereas VaR is not.

Minimising β -CVaR

We previously estimated $\min \mathbb{E}(h(x, \xi))$ by

$$\min_{x,s,f} \left\{ \sum_{t=1}^T \frac{1}{T} h(x, \xi)^{(t)} \right\}.$$

Minimising β -CVaR

We previously estimated $\min \mathbb{E}(h(x, \xi))$ by

$$\min_{x,s,f} \left\{ \sum_{t=1}^T \frac{1}{T} h(x, \xi)^{(t)} \right\}.$$

By a theorem from Rockafellar and Uryasev (2002) we have

$$\min \beta\text{-CVaR} = \min \left(\alpha + \frac{1}{1-\beta} \mathbb{E}(h(x, \xi) - \alpha)_+ \right),$$

Minimising β -CVaR

We previously estimated $\min \mathbb{E}(h(x, \xi))$ by

$$\min_{x,s,f} \left\{ \sum_{t=1}^T \frac{1}{T} h(x, \xi)^{(t)} \right\}.$$

By a theorem from Rockafellar and Uryasev (2002) we have

$$\min \beta\text{-CVaR} = \min \left(\alpha + \frac{1}{1-\beta} \mathbb{E}(h(x, \xi) - \alpha)_+ \right),$$

and so we can similarly estimate $\min \beta$ -CVaR by

$$\min_{\alpha,x,s,f} \left\{ \alpha + \frac{1}{1-\beta} \sum_{t=1}^T \frac{1}{T} \left(h(x, \xi)^{(t)} - \alpha \right)_+ \right\}.$$

Auxiliary Variables

Python cannot compute the minimum of a maximum, so we introduce a new variable z , a vector of length T . We then implement

$$\min_{\alpha, x, s, f, z} \left\{ \alpha + \frac{1}{1 - \beta} \sum_{t=1}^T \frac{1}{T} z^{(t)} \right\},$$

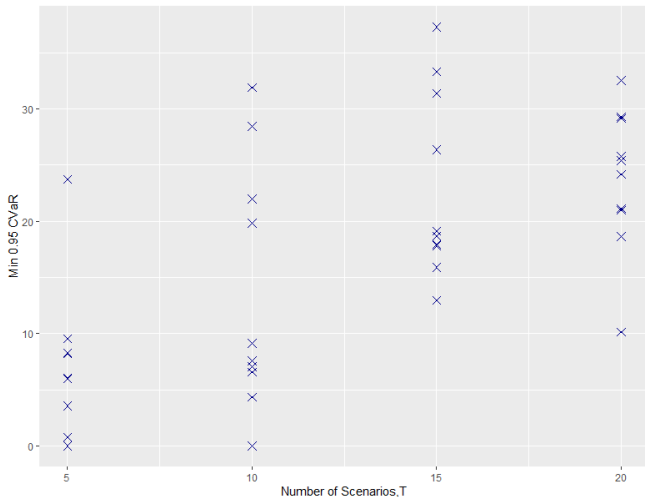
with additional constraints

$$\begin{aligned} z^{(t)} &\geq h(x, \xi)^{(t)} - \alpha, \\ z^{(t)} &\geq 0, \end{aligned}$$

for all $t \in \{1, \dots, T\}$.

Optimality Gap Replications

This shows the minimum 0.95-CVaR as the value of T changes.



Bibliography

- Acerbi, C. and Tasche, D. (2002). On the coherence of expected shortfall. *Journal of Banking & Finance*, 26(7):1487–1503.
- Chen, K. and Cheng, T. (2022). Measuring tail risks. *The Journal of Finance and Data Science*.
- Rockafellar, R. T. and Uryasev, S. (2002). Conditional value-at-risk for general loss distributions. *Journal of banking & finance*, 26(7):1443–1471.
- Sen, S., Doverspike, R. D., and Cosares, S. (1994). Network planning with random demand. *Telecommunication systems*, 3(1):11–30.

Thank You

Thank you for listening. Any questions?