EXTREMES WITH COVARIATE EFFECTS -PRACTICALITIES AND PROBLEMS

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Overview

Motivation

- Regulatory requirements ad-hoc (if not inconsistent) w.r.t. accommodation of covariate effects and estimation of (e.g.) directional and seasonal design values for coastal and ocean structures.
- Statistics literature provides a framework for rational and consistent estimation, but many issues unresolved.

This talk

- Analysis procedure with application to North Sea design
- Some pressing issues in modelling and interpretation

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Procedure in a nut shell: I

- Hindcast data for multiple locations in neighbourhood. Extract *storm peaks* (to eliminate temporal dependence) over threshold *u*.
- Assume extremal characteristics of all locations marginally identical, although dependent. Goal is to estimate distribution of *n*-year return value q_n for any single location.
- Estimation using NHPP: storm arrival rate $\mu \text{, GP}$ shape γ and scale $\sigma.$
- Accommodate covariate effects: μ , γ , σ and u vary with covariates (e.g. direction, season, time). u estimated before hand as high (local) quantile (sensitivity to threshold choice).
- Maximise likelihood, penalised by parameter roughness w.r.t. covariates. Diagnostics for model fit. Cross-validation for optimal roughness. Block bootstrap for parameter uncertainty pointwise.

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Procedure in a nut shell: II

- Simulate to estimate properties of q_n .
- Estimate q_n also for partitions w.r.t. covariates. Estimate and accommodate storm dissipation effects.
- Present findings in engineering terms.

Procedure published here:

- Effect of combining locations on estimation uncertainty (Jonathan and Ewans 2007b).
- Illustrations of extent of covariate effects on extreme quantile estimates (Jonathan et al. 2008).

Modelling directional extremes in the Gulf of Mexico and Northern North Sea (Jonathan and Ewans 2007a, Ewans and Jonathan 2008).

Data and pre-processing

- Significant wave height H_S values from a Northern North Sea hindcast, for October 1964 to September 1998 inclusive, at 3 hour intervals.
- For approximate location of 2° longitude, 61° latitude, selected 100 grid points on 10×10 rectangular lattice covering an area of approximately 5° longitude, 3° latitude centred at the location of interest.
- For each storm period for each grid point, isolated storm peak significant wave height, H_S^{sp} , and corresponding wave direction, θ .
- Estimated threshold for extreme value analysis based on local quantile (with direction)
- Estimated directional dissipation for a storm (w.r.t. storm peak H_S , with direction)

Pre-processing EDA (u, ρ)

Exploratory analysis



Quantiles of $H_S^{sp}(\theta)$



Conditional densities of $H_S^{sp}(\theta)$



 H_S^{sp} for NE and SW



 θ for NE and SW

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Jonathan & Ewans, UK Extremes 2008, Lancaster

Extremes with covariates

Threshold and dissipation



Threshold $u(\theta)$

- Given interval Θ of θ , threshold estimated empirically as quantile of $H_S^{sp}(\Theta)$
- Local median used for modelling henceforth



Dissipation $\rho(\Delta \theta)$

- For a given storm event, $\rho(\Delta\theta) = H_S(\Delta\theta)/H_S^{sp}$
- Dissipation quantifies influence of storm on extremes in directions other than θ

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Generalised Pareto Modelling: I

Given $\{X_i\}_{i=1}^n$, $\{\theta_i\}_{i=1}^n$, distribution of storm peaks above variable threshold $u(\theta)$ assumed GP with cdf $F_{X_i|\theta_i,u}$:

$$egin{array}{rcl} {\mathcal F}_{{\mathcal X}_i | heta_i, u} \left(x
ight) &=& {\mathcal P} \left({\mathcal X}_i \le x | heta_i, u \left(heta_i
ight)
ight) \ &=& 1 - \left({1 + rac{\gamma (heta_i)}{\sigma (heta_i)} \left({x - u \left({ heta_i}
ight)}
ight)
ight) _+ ^{ - rac{1}{\gamma (heta_i)} } \end{array}$$

 γ and σ vary smoothly with θ , assumed to follow Fourier form $\sum_{k=0}^{p}\sum_{b=1}^{2}A_{abk}t_{b}(k\theta).$

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Generalised Pareto Modelling: II

Penalised negative log likelihood is I^* :

$$I^* = \sum_{i=1}^n I_i + \lambda \left(R_\gamma + \frac{1}{w} R_\sigma \right)$$

Unpenalised negative log likelihood is:

$$I_{i} = \log \sigma \left(\theta_{i}\right) + \left(\frac{1}{\gamma\left(\theta_{i}\right)} + 1\right) \log \left(1 + \frac{\gamma\left(\theta_{i}\right)}{\sigma\left(\theta_{i}\right)}\left(X_{i} - u\left(\theta_{i}\right)\right)\right)_{+}$$

Roughness of γ is given by:

$$R_{\gamma} = \int_{0}^{2\pi} \left(\frac{\partial^{2}\gamma}{\partial\theta^{2}}\right)^{2} d\theta = \sum_{k=1}^{p} \pi k^{4} \left(\sum_{b=1}^{2} A_{1bk}^{2}\right)$$

Analogous expression for roughness of $\boldsymbol{\sigma}$

Cross-validation for roughness



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Forms of γ and σ with block bootstrap



 $\gamma(\theta), \sigma(\theta)$, with block bootstrap 95% confidence interval

Poisson Modelling: I

Non-homogeneous Poisson process model. The negative log-likelihood written:

$$I(\mu, \gamma, \sigma) = I_N(\mu) + I_W(\gamma, \sigma)$$

where I_N is the (negative) log-density of the total number of exceedances (with rate argument μ), and I_W is the (negative)log-conditional-density of exceedances given a known total number N). Inferences on μ made separately from those on γ and σ .

The Poisson process log-likelihood, for arrivals at times $\{t_i\}_{i=1}^n$ in period P_0 is:

$$I_N(\mu) = -(\sum_{i=1}^n \log \mu(t_i) - \int_{P_0} \mu(t) dt)$$

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Poisson Modelling: II

Or approximately (Chavez-Demoulin and Davison 2005):

$$\hat{l}_N(\mu) = -(\sum_{j=1}^m c_j \log \mu(j\delta) - \delta \sum_{j=1}^m \mu(j\delta))$$

where $\{c_j\}_{j=1}^m$ is the number of occurrences in each of the *m* sub-intervals. W

We estimate storm occurrence rate adopting a Fourier form for Poisson intensity μ as a function of θ , penalising its roughness R_{μ} :

$$\hat{l}_N^*(\mu) = \hat{l}_N(\mu) + \kappa R_\mu$$

 R_{μ} has form analogous to that of R_{γ} or R_{σ} . Cross-validation and block bootstrapping used similarly.

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Defining appropriate directional sectors



Optimal boundaries to reduce within-sector variability

100-year storm peak cdf using directional γ , σ



100-year storm peak cdf using constant γ , σ



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Design values for different design strategies

- Specification of an omni-directional non-exceedance probability does not uniquely define *sector* design values.
- Different strategies possible:
 - Design to omnidirectional H_{S100}^{sp}
 - Design to equal sector non-exceedence probability
 - Design to optimise a specified cost function

Sector	Angle	Risk-Cost Optimal			Omni-directional			Equal Non-exceedence probability		
Directional model		RiskCost	Value (m)	Quantile	RiskCost	Value (m)	Quantile	RiskCost	Value (m)	Quantile
S1	[20,130)	8.6	14.0	0.99	12.3	17.5	1.00	9.2	10.9	0.85
S2	[130,220)		15.6	0.82		17.5	0.94		15.8	0.85
S3	[220,270)		18.6	0.73		17.5	0.57		19.8	0.84
S4	[270,20)		16.6	0.84		17.5	0.92		16.6	0.84
Constant model		RiskCost	Value (m)	Quantile	RiskCost	Value (m)	Quantile	RiskCost	Value (m)	Quantile
S1	[20,130)	7.7	14.0	0.92	10.2	16.0	0.99	7.9	13.3	0.85
S2	[130,220)		16.2	0.82		16.0	0.79		16.3	0.85
S 3	[220,270)		15.9	0.81		16.0	0.83		16.0	0.84
S4	[270,20)		16.2	0.82		16.0	0.79		16.3	0.84

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Issues

Generic

- Sample size (c.f. estimates required for long return periods)
- Measurement (or hindcast) uncertainty (especially for extreme values)
- Temporal dependence (.:. "storm peak" analysis)
- Spatial dependence (... block bootstrap)
- Model form (e.g. GP versus Weibull ...) and complexity
- Transformation of variables (e.g. weighting locally w.r.t. covariate)
- Combination of variables (e.g. joint modelling, *structural response based* analysis)

Specific

- Reflecting model specification and fitting uncertainty in design values
 - Threshold selection
 - Model stiffness
 - Dissipation
- Specification and interpretation of design conditions in engineering context

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Model selection - a toy problem

- Two homogeneous directional sectors S_1 and S_2 , extremes are GP-distributed
- γ , σ and u values potentially different between sectors
- Random sample size 1250 from each sector corresponding to 25 years
- Test

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$$H_0$$
: $\gamma_1 = \gamma_2$, $\sigma_1 = \sigma_2$

- H_A : $\gamma_1 \neq \gamma_2$, $\sigma_1 \neq \sigma_2$
- Cases $(\gamma_1, \sigma_1, u_1)$ and $(\gamma_2, \sigma_2, u_2)$
 - Case 1 (-0.1,2,4) and (-0.3,4,4)
 - Case 2 (-0.1,3,4) and (-0.3,4,6)

Case 1 - Model selection



Sector densities (theory)



Probability of rejecting H_0

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Overview Data Modelling Design Issues Conclusions References

Overview Toy-Outline Toy-Case1 Toy-Case2 $q(\gamma)$ Seasonal

Case 1 - Parameter estimates with threshold



 H_A : S_1 estimates with threshold



 H_A : S_2 estimates with threshold

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H_0 : estimates with threshold

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Case 1 - Estimated median 100-year maximum



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Case 1 - Implications for long return periods



Median P-year event: true (dot), H_A (dash),

 H_0 (full, for thresholds 6,7,...,11m)

Case 2 - Model selection



Sector densities (theory)



Probability of rejecting H_0

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Overview Data Modelling Design Issues Conclusions References

Overview Toy-Outline Toy-Case1 Toy-Case2 $q(\gamma)$ Seasonal

Case 2 - Parameter estimates with threshold



 H_A : S_1 estimates with threshold



 H_A : S_2 estimates with threshold

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H_0 : estimates with threshold

Case 2 - Estimating median 100-year maximum



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Interpretation: Ratios of extreme quantiles with γ



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Interpretation: seasonal extremes



Design values for short-term deployments in Gulf of Mexico

Future

Future work

- Model multiple covariate effects (e.g. more general smoothers).
- Model spatial and temporal dependence explicitly (e.g. extreme quantiles for region rather than single location).
- Improved modelling of dissipation effects.
- Jointly model multiple variables (wind, waves, current, e.g. Heffernan and Tawn 2004), compare inferences with *response-based* approaches.
- Extremes estimates incorporating uncertainties from model and threshold specification (e.g. predictive distributions).
- Models that better exploit the underlying physics (e.g. for hurricanes)
- Influence design practice. Regulators (e.g. API) currently reviewing methods for seasonal and directional design. Bridge industry and academia, communicate.

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