

# On Extremal Behavior of Gaussian Chaos

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Let  $\xi = (\xi_1, \xi_2, \dots, \xi_d)$  be a normally distributed random vector in  $\mathbb{R}^d$  with zero mean and covariance matrix  $B$ ,  $B_{ij} := \mathbb{E}\xi_i\xi_j$ . A problem of great interest is to analyze the asymptotic behavior of the distribution tail

of the product  $\prod_{i=1}^d \xi_i$ . This problem arises in various domains, for example in stochastic geometry, random difference equations, and risk theory.

Consider a more general case of functions of the vector  $\xi$ , namely, the so-called Gaussian chaos  $h(\xi)$ , where  $h: \mathbb{R}^d \rightarrow \mathbb{R}$  is a continuous homogeneous function of order  $\alpha > 0$ ; i.e.,  $h(x\mathbf{t}) = x^\alpha h(\mathbf{t})$  for all  $x > 0$  and  $\mathbf{t} \in \mathbb{R}^d$ . Traditionally, in the literature, the term Gaussian chaos of order  $\alpha \in \mathbb{N}$  is referred to the case where  $g$  is a homogeneous polynomial of degree  $\alpha$ . This concept goes back to Wiener [14], who was the first to consider processes of polynomial chaos. We follow a broader treatment of the concept of Gaussian chaos.

The distribution of  $\xi$  is equal to the distribution of  $\sqrt{B}\boldsymbol{\eta}$  if the vector  $\boldsymbol{\eta} = (\eta_1, \eta_2, \dots, \eta_d)$  has independent coordinates with a standard normal distribution. Then

$$\begin{aligned} \mathbb{P}\{h(\xi) > x\} &= \mathbb{P}\{h(\sqrt{B}\boldsymbol{\eta}) > x\} \\ &= \mathbb{P}\{g(\boldsymbol{\eta}) > x\}, \end{aligned}$$

where  $g(\mathbf{u}) = h(\sqrt{B}\mathbf{u})$ . The continuous function  $g: \mathbb{R}^d \rightarrow \mathbb{R}$  is also homogeneous of order  $\alpha$  like  $h$ . Thus, the problem is reduced to the case of a unit covariance matrix. For this reason, in what follows, we study  $g(\boldsymbol{\eta})$ . By virtue of homogeneity,

$$\begin{aligned} \mathbb{P}\{g(\boldsymbol{\eta}) > x\} &= \mathbb{P}\{(g(x^{-1/\alpha}\boldsymbol{\eta}) > 1\} \\ &= \frac{x^{d/\alpha}}{(2\pi)^{d/2}} \int_{\{\mathbf{v}: g(\mathbf{v}) > 1\}} e^{-x^{2/\alpha}|\mathbf{v}|^2/2} d\mathbf{v}. \end{aligned} \quad (1)$$

Therefore, the asymptotic behavior of probability (1) can be determined using a version of the Laplace asymptotic method (see, for example, [3]). Define

$$\begin{aligned} c^2 &:= \min\{|\mathbf{u}|^2: g(\mathbf{u}) \geq 1\} \\ &= \min\{|\mathbf{u}|^2: g(\mathbf{u}) = 1\}, \end{aligned}$$

where the last equality follows from the homogeneity of  $g$ . Since  $g$  is continuous, we have  $c^2 > 0$ . To apply the Laplace method, we consider the set

$$\begin{aligned} \mathcal{C} &:= \arg \min\{|\mathbf{u}|: g(\mathbf{u}) = 1\} \\ &= \{\mathbf{u}: |\mathbf{u}| = c \text{ and } g(\mathbf{u}) = 1\}, \end{aligned}$$

which lies on a sphere of radius  $c$ . Assume that this set is a smooth finitely connected manifold of dimension  $r$  and the structure of the function  $g$  near this manifold is typical of the Laplace method. Define  $g(\boldsymbol{\varphi}) := g(\mathbf{u}/|\mathbf{u}|)$ , where  $\boldsymbol{\varphi} = (\varphi_1, \varphi_2, \dots, \varphi_{d-1}) \in \Pi := [0, \pi]^{d-2} \times [0, 2\pi)$  are the spherical coordinates of the vector  $\mathbf{u}/|\mathbf{u}|$  on the unit sphere  $S_{d-1}$ . The manifold on the parallel-epiped  $\Pi$  that corresponds to  $\mathcal{C}$  is denoted by  $\mathcal{C}_\varphi$ . The Jacobian of the transition to spherical coordinates in  $\mathbb{R}^d$  is designated as  $J(r, \boldsymbol{\varphi})$ . Let  $g''(\boldsymbol{\varphi})$  denote the Hessian of a function  $g(\boldsymbol{\varphi})$ , and let  $\lambda(A)$  stand for the smallest (in absolute value) nonzero eigenvalue of a symmetric matrix  $A$ .

**Theorem 1.** *Let  $g: \mathbb{R}^d \rightarrow \mathbb{R}$  be a continuous homogeneous function of order  $\alpha > 0$ , and let  $\dim \mathcal{C}_\varphi = r \in [0, d-1]$ . If the corresponding function  $g(\boldsymbol{\varphi}): \Pi \rightarrow \mathbb{R}$  is three times differentiable and*

$$\text{rank} g''(\boldsymbol{\varphi}) \equiv d-1-r, \quad \inf_{\boldsymbol{\varphi} \in \mathcal{C}_\varphi} \lambda(g''(\boldsymbol{\varphi})) > 0$$

(the Hessian is uniformly nonsingular on  $\mathcal{C}_\varphi$ ), then

$$\begin{aligned} \mathbb{P}\{g(\boldsymbol{\eta}) > x\} &= \mathcal{H} x^{(r-1)/\alpha} e^{-c^2 x^{2/\alpha}/2} (1 + O(x^{-2/\alpha})) \quad (2) \\ &\text{as } x \rightarrow \infty, \end{aligned}$$

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$$\mathcal{H} := \frac{1}{(2\pi)^{(r+1)/2}} \frac{\alpha^{(d-1-r)/2}}{c^{1-r+\alpha(d-1-r)/2}} \times \int_{\mathcal{C}_\varphi} \frac{J(1, \varphi)}{\sqrt{|\det g''_{d-1-r}(\varphi)|}} dV_\varphi,$$

where  $dV_\varphi$  is the volume element of the manifold  $\mathcal{C}_\varphi \subset \Pi$  and  $\det g''_{d-1-r}(\varphi)$  is any nonzero minor of the Hessian  $g''(\varphi)$  of order  $d - 1 - r$ . Relation (2) can be differentiated, which gives asymptotics of the distribution density of the Gaussian chaos  $g(\boldsymbol{\eta})$ .

Note that, as in the classical case of the Laplace method [3], assuming that  $g$  has higher smoothness, we can obtain asymptotic expansions of the considered probability and density in powers of  $x$ . In the case  $r = 0$ , i.e., when  $\mathcal{C} = \{\mathbf{t}_1, \mathbf{t}_2, \dots, \mathbf{t}_k\}$ , where  $\mathbf{t}_i$  are isolated absolute minimizers of  $g$  in the integration domain, the theorem is proved by directly applying Theorem 4.2 from [3]. The integral in the expression for  $\mathcal{H}$  becomes a sum over the points  $\varphi_i \in \Pi$  corresponding to the points  $\mathbf{t}_i$ . In the general case, we apply a version of the Laplace method for parameter-dependent functions, which are used to prove the possibility of integration. On each map of an atlas with sufficiently small maps on the manifold  $\mathcal{C}_\varphi$ , we construct a coordinate system with the first  $r$  coordinates being parameters. When they are fixed, the minimum of the amplitude (of the argument of the exponential) is reached at a unique point of a neighborhood of the map. Next, the standard Laplace method is applied and the maps of the atlas are integrated with respect to these parameters on all neighborhoods in  $\Pi$ .

By Theorem 1, the Gaussian chaos is a subexponential random variable if  $\alpha > 2$ . The subexponentiality of random variables is an important concept in various applications (see, for example, [4]). The Gaussian chaos is subexponential under rather weak constraints on the function  $h$ . For example, let  $h$  be nonnegative. The  $d$ -dimensional centered Gaussian vector  $\boldsymbol{\eta}$  with a unit covariance matrix can be represented as the product  $\boldsymbol{\eta} \stackrel{d}{=} \chi \boldsymbol{\mu}$  of independent values  $\chi$  and  $\boldsymbol{\mu}$ , where  $\chi^2 = \sum_{i=1}^d \eta_i^2$  has a chi-square distribution  $\chi^2$  with  $d$  degrees of freedom, while  $\boldsymbol{\mu}$  has a uniform distribution on the unit sphere  $S_{d-1} \subset \mathbb{R}^d$ . The Gaussian random vector  $\boldsymbol{\xi} = \sqrt{B}\boldsymbol{\eta} = \chi\sqrt{B}\boldsymbol{\mu}$  has the covariance matrix  $B$ . Therefore, since  $h$  is homogeneous for any  $x > 0$ , we have

$$\mathbb{P}\{h(\boldsymbol{\xi}) > x\} = \mathbb{P}\{\chi^\alpha h(\sqrt{B}\boldsymbol{\mu}) > x\}. \tag{3}$$

If  $h(\sqrt{B}\boldsymbol{\mu})$  is a positive bounded random variable, then, according to [2, Corollary 2.5], the random vari-

able  $h(\boldsymbol{\xi})$  is subexponential for  $\alpha > 2$ , because the distribution  $\chi^\alpha$  then has a Weibull type density

$$\frac{1}{\alpha \cdot 2^{d/2-1} \Gamma(d/2)} x^{d/\alpha-1} e^{-x^{2/\alpha}/2}$$

with  $2/\alpha < 1$ , which means subexponentiality.

It follows from (3) that, if  $h$  is bounded on the unit sphere  $S_{d-1}$ , i.e.,  $h^* := \max\{h(\mathbf{u}) : |\mathbf{u}| = 1\} < \infty$ , then estimates

$$\mathbb{P}\{h(\boldsymbol{\xi}) > x\} \leq \mathbb{P}\{\chi^\alpha > x/h^*\} \leq \frac{1}{\alpha \cdot 2^{d/2-1} \Gamma(d/2)} \int_{x/h^*}^\infty y^{d/\alpha-1} e^{-y^{2/\alpha}/2} dy.$$

This explicit upper bound improves the one obtained in [10, Corollary 1]. In our conditions, it is better than the bound that can be derived from [1, Theorem 4.3].

Theorem 1 underlies a unified approach to different problems. Below are some examples.

**Example 1.** (Product of independent  $N(0, 1)$  random variables) Let  $\boldsymbol{\eta} = (\eta_1, \eta_2, \dots, \eta_d)$  be a standard Gaussian vector and  $g(\mathbf{u}) = u_1 u_2 \dots u_d$ . We have  $\alpha = d$ ,  $c^2 = d$ , and  $\mathcal{C} = \{(\pm 1, \dots, \pm 1)$  with an even number of negative coordinates} consists of  $2^{d-1}$  points. Applying Theorem 1 yields the asymptotics

$$p_{\eta_1 \dots \eta_d}(x) = \frac{2^{(d-1)/2}}{\sqrt{2\pi d}} x^{1/d-1} e^{-dx^{2/d}/2} (1 + O(x^{-2/d}))$$

as  $x \rightarrow \infty$ .

This asymptotic relation can be intuitively interpreted as follows (see, e.g., [13]): the product takes the most probable large value when all the multipliers are roughly identical; therefore,  $p_{\eta_1 \dots \eta_d}(x)$  asymptotically resembles the product of  $d$  densities at the same point  $x^{1/d}$ .

For the product of the coordinates of an arbitrary Gaussian vector  $\boldsymbol{\xi}$  with a covariance matrix  $B$ , we have a similar formula based on the representation  $\boldsymbol{\xi} = \sqrt{B}\boldsymbol{\eta}$ , but the computation of the constants encounters certain difficulties.

**Example 2.** (Quadratic forms of independent  $N(0, 1)$  random variables.) Let  $g(\boldsymbol{\eta}) = \sum_{i=1}^d a_i \eta_i^2$ , where the constants  $a_i \in \mathbb{R}$  are such that  $a_1 \leq a_2 \leq \dots \leq a_{d-r} < a_{d-r+1} = \dots = a_d = a, a > 0$ .

Since

$$g(\mathbf{u}) = \sum_{i=1}^{d-r} a_i u_i^2 + a \sum_{i=d-r+1}^d u_i^2$$

and  $a_i < a$  for  $i \leq d - r$ , the minimum of  $|\mathbf{u}|^2$  on the set  $g(\mathbf{u}) = 1$  is reached at points  $\mathbf{u}$  satisfying  $u_{d-r+1}^2 + \dots + u_d^2 = \frac{1}{a}$  and  $u_1 = u_2 = \dots = u_{d-r} = 0$ , so that  $c^2 = \frac{1}{a}$ . If

$r = 1$ , the set  $\mathcal{C}_\varphi$  consists of two points  $(\frac{\pi}{2}, \dots, \frac{\pi}{2}, \frac{\pi}{2})$  and  $(\frac{\pi}{2}, \dots, \frac{\pi}{2}, \frac{3\pi}{2})$ . By using Theorem 1, we can find that

$$\mathbb{P}\left\{\sum_{i=1}^d a_i \eta_i^2 > x\right\} = \frac{1}{2^{r/2-1} \Gamma(r/2)} \prod_{i=1}^{d-r} \frac{1}{\sqrt{1-a_i/a}} (x/a)^{r/2-1} e^{-x/2a} (1 + O(1/x))$$

as  $x \rightarrow \infty$ , which agrees (up to the first-order asymptotics) with the results of [6] (see also [11, 12] or [7, Theorem 1]). This also supplements the upper bounds obtained in [5, 9].

**Example 3.** (Scalar product) The quadratic forms in Example 2 are closely related to  $g(\boldsymbol{\eta}, \boldsymbol{\eta}^*) = \sum_{i=1}^d a_i \eta_i \eta_i^*$ , where  $\eta_i$  and  $\eta_i^*$ ,  $i \leq d$ , are independent  $N(0, 1)$  random variables and  $a_i \in \mathbb{R}^+$ . Indeed, since  $\eta_i \eta_i^*$  coincides in distribution with

$$\frac{\eta_i + \eta_i^*}{\sqrt{2}} \frac{\eta_i - \eta_i^*}{\sqrt{2}} = \frac{\eta_i^2 - \eta_i^{*2}}{2},$$

we have the distribution equality

$$g(\boldsymbol{\eta}, \boldsymbol{\eta}^*) \stackrel{d}{=} \frac{1}{2} \left( \sum_{i=1}^d a_i \eta_i^2 - \sum_{i=1}^d a_i \eta_i^{*2} \right),$$

and, to the quadratic form on the right, we can apply the result of Example 2, with the dimension replaced by  $2d$  and with the parameter  $r$  replaced by the number of maximal  $a_i$ . Some results for scalar products can be found in [8].

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